

# On A Law of Large Number for Last Passage Percolation on Complete Graph

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- **Introduction**
- **Results**
- **On the proofs**

## Last passage percolation on complete graph

$G_n = ([n], E_n)$ : the complete graph, where

$$[n] = \{1, 2, \dots, n\}, E_n = \{\langle i, j \rangle : 1 \leq i < j \leq n\}.$$

$\{X_e : e \in E_n\}$ : i.i.d. **positive** random variables, edge passage times.

$\Pi_{1,n}$ : the set of all self-avoiding paths between vertex 1 and  $n$ .

$T(\pi)$ : the passage time of path  $\pi \in \Pi_{1,n}$

$$T(\pi) = \sum_{e \in \pi} X_e.$$

$W_n$ : the largest passage time among all self-avoiding paths from 1 to  $n$ , i.e.,

$$W_n = \sup_{\pi \in \Pi_{1,n}} T(\pi).$$

# 1. The time constant $\mu$ exists!

## Theorem 1

*For any distribution of edge passage time, the time constant of the model exists. More precisely, there exists some constant  $0 < \mu \leq \infty$ , such that*

$$\frac{W_n}{n} \rightarrow \mu \text{ a.s.}$$

*as  $n \rightarrow \infty$ . In particular, when  $\mu < \infty$ , the above convergence is also in  $L_1$ . Furthermore,  $\mu$  coincides with the **essential supremum** of  $X_e$ , i.e.*

$$\mu = \inf\{x : \mathbb{P}(X_e > x) = 0\}.$$

## 2. Lower and upper bounds for $W_n/n$

Let  $F(x) = \mathbb{P}(X_e \leq x)$  and  $H(x) = 1 - F(x)$ ,  $x \in \mathbb{R}$  be the distribution function and the tail probability function of  $X_e$ . By Theorem 1, when  $\mu = \infty$ ,  $W_n/n$  tends to  $\infty$  as  $n \rightarrow \infty$ . Here, we give lower and upper bounds to  $W_n/n$  as in the following theorem.

### Theorem 2

Suppose that  $\mu = \infty$ . Let  $f$  and  $g$  be two functions such that  $H(f(n)) = \ln n/n$  and  $n^2 H(g(n)) \rightarrow 0$  as  $n \rightarrow \infty$ . Then, for any  $\epsilon > 0$

$$\lim_{n \rightarrow \infty} \mathbb{P} \left( (1 - \epsilon)f(n) \leq \frac{W_n}{n} \leq g(n) \right) = 1.$$

# A classification of increasing function

## Typically increasing function:

Let's consider the usual exponential function  $\exp(x)$ . For any  $n \in \mathbb{Z}$ , write  $\exp^n(x)$  as the  $n$ -th iteration of the exponential function. Note that, by convention,  $\exp^0(x) = x$ ,  $\exp^{-1}(x) = \log(x)$ , and  $\exp^{-n}(x)$  is the  $n$ -th iteration of  $\log(x)$  for  $n > 0$ .

A function  $\psi$  is called *typically increasing*, if  $\psi$  has the form

$$\psi(x) = a[\exp^n(x)]^b + c$$

with  $a, b > 0$ ,  $c \in \mathbb{R}$ ,  $n \in \mathbb{Z}$ , for large enough  $x \in \mathbb{R}$ .

## Regularly increasing function:

A strictly increasing function  $\phi$  is called *regularly increasing* if, it is differentiable (at all large enough  $x$ ), and increases “regularly” compared to typically increasing functions:

for any *typically increasing* function  $\psi$ , **IF** for some increasing sequence  $\{x_n\}$  with  $\lim_{n \rightarrow \infty} x_n = \infty$ , one has  $\phi(x_n) \geq \psi(x_n)$ , **THEN**, for all large enough  $x$ ,

- $\phi(x) \geq \psi(x)$ ;
- $\phi(x) - \psi(x)$  and  $\psi^{-1}(x) - \phi^{-1}(x)$  are monotone in  $x$ .

Where  $\phi^{-1}$  and  $\psi^{-1}$  are the inverses of  $\phi$  and  $\psi$ .

“Regular” means:  $\phi - \psi$  and  $\psi^{-1} - \phi^{-1}$  do not change their monotonicity **frequently!**



# Assumptions

Rewrite  $H(x) = \mathbb{P}(X_e > x)$ , the tail probability of  $X_e$ , in the form

$$H(x) = e^{-\beta(x)}.$$

We assume  $\beta(x)$  satisfy the following **technical conditions**:

**A1:**  $\beta(x)$  is second-order differentiable and  $\beta''(x)$  is continuous;

**A2:**  $\gamma(x) := \beta'(x)$  is monotone. In the case when  $\lim_{x \rightarrow \infty} \gamma(x) = +\infty$ ,  $\gamma(x)$  increases strictly and the function  $s(x)$  solving the equation

$$\gamma(s(x)) = D\gamma(x), \quad D > 1,$$

is **regularly increasing**.

# Statement of Result

## Theorem 3: Law of Large Number

Suppose that  $H(x) = \mathbb{P}(X_e > x) = e^{-\beta(x)}$ . If for some  $t > 0$ ,

$$\varphi(t) := \mathbb{E}(e^{tX_e}) < \infty,$$

and the above assumptions A1 and A2 hold, then

$$\frac{W_n}{n\beta^{-1}(\log n)} \rightarrow 1$$

in probability as  $n \rightarrow \infty$ . Where  $\beta^{-1}$  is the inverse of  $\beta$ .

For any  $x > 0$ , let  $I(x)$  be the *Legendre transform* of the cumulant generating function  $\log \varphi$ , namely,

$$I(x) = \sup_{t>0} [xt - \log \varphi(t)].$$

First of all, we have the following lemma.

**Lemma 1:** Under the condition of Theorem 3, for any  $\epsilon > 0$ , one has

$$I[(1 + \epsilon)\beta^{-1}(\log n)] \geq \log n$$

for all large enough  $n$ .

## Slowly varying function

A function  $f$  is called *slowly varying*, if for any  $0 < c < 1$ ,

$$\lim_{x \rightarrow +\infty} \frac{f(cx)}{f(x)} = 1.$$

Let

$$\alpha(x) := \frac{\beta(x)}{x},$$

then by the condition of Theorem 3

$$\lim_{x \rightarrow \infty} \alpha(x) = \alpha_0 \in (0, +\infty]$$

and when  $\alpha_0 < +\infty$ , the function  $\alpha$  is *slowly varying*.

# Proof of Lemma 1

We prove Lemma 1 in two steps.

**Step 1:** The case  $\alpha(x) = \frac{\beta(x)}{x}$  is slowly varying.

$$\varphi(t) = \int_0^{+\infty} \mathbb{P}(e^{tX_e} > x) dx = 1 + \int_1^{+\infty} H\left(\frac{\ln x}{t}\right) dx.$$

For large enough  $x$  and small  $c > 0$ , let  $\lambda_0 = \alpha(c\beta^{-1}(x))$ , one has

$$\begin{aligned} \varphi(t) &\leq e^{c\beta^{-1}(x)t} + \int_{e^{c\beta^{-1}(x)t}}^{+\infty} e^{-\frac{\lambda_0}{t} \log y} dy \\ &= e^{c\beta^{-1}(x)t} \left( \frac{\lambda_0 - \left(1 - e^{-c\beta^{-1}(x)\lambda_0}\right)t}{\lambda_0 - t} \right) \end{aligned}$$

for all  $0 < t < \lambda_0$ .

Then

$$\begin{aligned} & I((1 + \epsilon)\beta^{-1}(x)) \\ & \geq \sup_{0 < t < \lambda_0} \left[ (1 + \epsilon - c)\beta^{-1}(x)t - \log \left( \frac{\lambda_0 - (1 - e^{-c\beta^{-1}(x)\lambda_0})t}{\log(\lambda_0 - t)} \right) \right] \\ & = \sup_{0 < t < \lambda_0} \left[ \frac{(1 + \epsilon - c)}{\alpha(\beta^{-1}(x))}xt - \log \left( \frac{\lambda_0 - (1 - e^{-c\beta^{-1}(x)\lambda_0})t}{\log(\lambda_0 - t)} \right) \right] \\ & = \frac{(1 + \epsilon - c)}{\alpha(\beta^{-1}(x))}xt_0 - \log \left[ \lambda_0 - (1 - e^{-c\beta^{-1}(x)\lambda_0})t_0 \right] + \log(\lambda_0 - t_0). \end{aligned}$$

Where  $t_0 = t_0(x)$  fits the supremum. It is straightforward to check that

$$\frac{\lambda_0}{t_0(x)} \rightarrow 1,$$

$$\frac{1}{x} \log \left[ \lambda_0 - \left( 1 - e^{-c\beta^{-1}(x)\lambda_0} \right) t_0 \right] \rightarrow 0 \quad \text{and}$$

$$\frac{1}{x} \log(\lambda_0 - t_0) \rightarrow 0$$

as  $x \rightarrow +\infty$ . Thus, for small enough  $c$ , we have

$$\frac{(1 + \epsilon - c)}{\alpha(\beta^{-1}(x))} t_0 \rightarrow (1 + \epsilon - c) \frac{\alpha(c\beta^{-1}(x))}{\alpha(\beta^{-1}(x))} > 1$$

for large enough  $x$ , then  $I((1 + \epsilon)\beta^{-1}(x)) \geq x$  for large enough  $x$ .

**Step 2:** The case  $\alpha(x) := \frac{\beta(x)}{x} \nearrow \infty$  and is **NOT** slowly varying.

In this step, for any  $\epsilon > 0$ , we try to find some  $t_0 = t_0(n)$ , such that

$$(1 + \epsilon)\beta^{-1}(\log n)t_0 - \log \varphi(t_0) \geq \log n, \text{ for all large } n.$$

Let  $\zeta(n) = \gamma(\beta^{-1}(\log n))/\alpha(\beta^{-1}(\log n))$  and let

$$t_0 = \gamma(\beta^{-1}(\log n)) = \zeta(n)\alpha(\beta^{-1}(\log n)).$$

Using the variable substitution  $s = \frac{\log x}{\zeta(n)\log n}$ , one has

$$\int_1^{+\infty} H\left(\frac{\log x}{t_0}\right) dx = \zeta(n) \log n \int_0^{+\infty} n^{g(s)} ds,$$

where

$$g(s) = \zeta(n)s - \frac{\beta(s\beta^{-1}(\log n))}{\log n}.$$



Then, for any  $s_1 > 0$ ,

$$\varphi(t_0) = 1 + \zeta(n) \log n \left[ \int_0^{s_1} n^{g(s)} ds + \int_{s_1}^{+\infty} n^{g(s)} ds \right]$$

Let  $s_1$  be the solution of equation

$$g(s) = -s$$

such that, for  $s > s_1$ ,

$$g(s) < -s.$$

Noticing that  $\beta(x)$  increases strictly, one has  $g'(1) = 0$ ,  $g'(s) > 0$  when  $s < 1$  and  $g'(s) < 0$  when  $s > 1$ , then

$$g(1) = \sup_{s \geq 0} g(s) = \zeta(n) - 1.$$

So,

$$\begin{aligned}\varphi(t_0) &= 1 + \zeta(n) \log n \left[ \int_0^{s_1} n^{g(s)} ds + \int_{s_1}^{+\infty} n^{g(s)} ds \right] \\ &\leq 1 + \zeta(n) \log n \left[ s_1 n^{g(1)} + \int_0^{+\infty} n^{-s} ds \right] \\ &\leq 1 + \zeta(n) \log n \cdot (s_1 n^{\zeta(n)-1} + \text{constant}).\end{aligned}$$

Now, for  $s_1$ , we **claim** that,

$$\forall b > 0, s_1(n) \leq n^b.$$

Then, for any  $b > 0$ ,

$$\begin{aligned}\varphi(t_0) &\leq 1 + \zeta(n) \log n \cdot (s_1 n^{\zeta(n)-1} + \text{constant}) \\ &\leq \zeta(n) \log n \cdot n^{\zeta(n)-1+b}\end{aligned}$$

for large enough  $n$ .

## About the claim

Actually, the equation  $g(s) = -s$  can be rewritten as

$$\gamma(\beta^{-1}(\log n)) + \alpha(\beta^{-1}(\log n)) = \alpha(s\beta^{-1}(\log n)).$$

Then

$$\gamma\left(\frac{s_1}{2}\beta^{-1}(\log n)\right) \leq 4\gamma(\beta^{-1}(\log n)).$$

Let  $s(x)$  be the root of the equation

$$\gamma(s(x)) = 4\gamma(x), x > 0. \quad (*)$$

By the following [Lemma 2](#), for any  $b > 0$ ,

$$s(x) \leq e^{bx}$$

for large  $x$ . Then,

$$s_1(n) \leq \frac{s_1(n)}{2} \beta^{-1}(\log n) \leq s(\beta^{-1}(\log n)) \leq s(\log n) \leq n^b$$

for large enough  $n$ .

## Lemma 2 and its proof

**Lemma 2:** Suppose that  $\lim_{x \rightarrow +\infty} \alpha(x) = +\infty$  and  $\alpha(x)$  is not slowly varying, let  $s(x)$  be the solution of equation (\*). Then under the condition of Theorem 3, one has

$$\forall b > 0, \quad s(x) \leq e^{bx}$$

for all large  $x$ .

**Proof:**

- 1  $s(x)$  is regularly increasing (by Assumption A2).
- 2 If not,  $\alpha$  is slowly varying.

Now, we have, for any  $b > 0$ ,

$$\begin{aligned}\varphi(t_0) &= 1 + \zeta(n) \log n \left[ \int_0^{s_1} n^{g(s)} ds + \int_{s_1}^{+\infty} n^{g(s)} ds \right] \\ &\leq 1 + \zeta(n) \log n \left[ s_1 n^{g(1)} + \int_0^{+\infty} n^{-s} ds \right] \\ &\leq \zeta(n) \log n \cdot n^{\zeta(n)-1+b}\end{aligned}$$

for large enough  $n$ .

Hence, we have, for any  $b > 0$

$$\begin{aligned}(1 + \epsilon)\beta^{-1}(\log n)t_0 - \log \varphi(t_0) &\geq (1 + \epsilon)\zeta(n) \log n - \log \varphi(t_0) \\ &\geq \left[1 + \epsilon\zeta(n) - b - \frac{\log \zeta(n) + \log \log n}{\log n}\right] \log n\end{aligned}$$

Noticing that  $\zeta(n) \geq 1$ , then, by taking  $b = \epsilon/2$ , we have

$$\begin{aligned}I((1 + \epsilon)\beta^{-1}(\log n)) &\geq \left[1 + \frac{1}{2}\epsilon\zeta(n) - \frac{\log \zeta(n) + \log \log n}{\log n}\right] \log n \\ &\geq \log n\end{aligned}$$

for large  $n$ . Thus, we finish the second step of the proof.

## Proof of Theorem 3

Lower bound part:

Theorem 2 has already given the lower bound

$$f(n) = \beta^{-1}(\log n - \log \log n),$$

so it suffices for us to prove that

$$\lim_{n \rightarrow \infty} \frac{\beta^{-1}(\log n - \log \log n)}{\beta^{-1}(\log n)} = 1. \quad (**)$$



When  $\alpha$  is *slowly varying*, then for any  $0 < c < 1$ , one has

$$1 \geq \frac{\beta^{-1}(\log n - \log \log n)}{\beta^{-1}(\log n)} \geq \frac{\beta^{-1}(c \log n)}{\beta^{-1}(\log n)} \rightarrow c$$

as  $n \rightarrow \infty$ , (\*\*) follows by taking  $c \nearrow 1$ .

Otherwise, one has  $\alpha_0 = +\infty$  and by A2,  $\gamma(x)$  is strictly increasing, hence,

$$\begin{aligned} & \beta^{-1}(\log n) - \beta^{-1}(\log n - \log \log n) \\ &= \int_{\log n - \log \log n}^{\log n} (\beta^{-1})'(y) dy \\ &= \int_{\log n - \log \log n}^{\log n} \frac{1}{\gamma(\beta^{-1}(y))} dy \\ &\leq \frac{\log \log n}{\gamma(\beta^{-1}(\log n - \log \log n))}. \end{aligned}$$

Then,

$$\begin{aligned} & \frac{\beta^{-1}(\log n) - \beta^{-1}(\log n - \log \log n)}{\beta^{-1}(\log n)} \\ & \leq \frac{\log \log n}{\log n - \log \log n} \frac{\alpha(\beta^{-1}(\log n - \log \log n))}{\gamma(\beta^{-1}(\log n - \log \log n))} \\ & \leq \frac{\log \log n}{\log n - \log \log n} \rightarrow 0 \end{aligned}$$

as  $n \rightarrow \infty$ . Thus we get (\*\*).

## Upper bound part:

We have

$$\begin{aligned} & \mathbb{P}(W_n \geq (1 + \epsilon)n\beta^{-1}(\log n)) \\ & \leq \sum_{\pi \in \Pi_{1,n}} \mathbb{P}(T(\pi) \geq (1 + \epsilon)n\beta^{-1}(\log n)) \quad (***) \\ & \leq n! \cdot \mathbb{P}(S_n \geq (1 + \epsilon)n\beta^{-1}(\log n)), \end{aligned}$$

where  $S_n = \sum_{i=1}^n X_i$  and  $\{X_i\}$  be i.i.d. random variables with the same distribution as  $X_e$ .

By Stirling's formula

$$n! = n^n e^{-n} \sqrt{2\pi n} \left(1 + O\left(\frac{1}{n}\right)\right)$$

and the basic large deviation inequality

$$\mathbb{P}(S_n \geq an) \leq e^{-I(a)n}, \text{ for } a > \mathbb{E}X_e,$$

the upper bound part of the theorem follows from (\*\*\*) and Lemma 1:

$$\begin{aligned} & \mathbb{P}(W_n \geq (1 + \epsilon)n\beta^{-1}(\log n)) \\ & \leq n! \cdot \mathbb{P}(S_n \geq (1 + \epsilon)n\beta^{-1}(\log n)), \\ & \leq n! \cdot e^{-nI((1+\epsilon)\beta^{-1}(\log n))} \\ & \leq e^{-n} \sqrt{2\pi n} \left(1 + O\left(\frac{1}{n}\right)\right) n^n e^{-n \log n} \\ & = e^{-n} \sqrt{2\pi n} \left(1 + O\left(\frac{1}{n}\right)\right) \rightarrow 0. \end{aligned}$$

**Thanks for your attention!**